# 8. Sampling \& Estimation 

Spring 2021
Matthew Blackwell

Gov 2002 (Harvard)

## Where are we? Where are we going?

- Last few weeks: probability, learning how to think about r.v.s


## Where are we? Where are we going?

- Last few weeks: probability, learning how to think about r.v.s
- Now: how to estimate features of underlying distributions with data.


## Where are we? Where are we going?

- Last few weeks: probability, learning how to think about r.v.s
- Now: how to estimate features of underlying distributions with data.
- How do we construct estimators? What are their properties?

1/ Point Estimation

## Motivating example

- Gerber, Green, and Larimer (APSR, 2008)


## Dear Registered Voter:

WHAT IF YOUR NEIGHBORS KNEW WHETHER YOU VOTED?
Why do so many people fail to vote? We've been talking about the problem for years, but it only seems to get worse. This year, we're taking a new approach. We're sending this mailing to you and your neighbors to publicize who does and does not vote.

The chart shows the names of some of your neighbors, showing which have voted in the past. After the August 8 election, we intend to mail an updated chart. You and your neighbors will all know who voted and who did not.

DO YOUR CIVIC DUTY - VOTE!

| MAPLE DR | Aug 04 | Nov 04 | Aug 06 |
| :--- | :--- | :--- | :--- |
| 9995 JOSEPH JAMES SMITH | Voted | Voted |  |
| 9995 JENNIFER KAY SMITH |  | Voted | - |
| 9997 RICHARD B JACKSON |  | Voted | - |
| 9999 KATHY MARIE JACKSON |  | Voted | - |

## Motivating Example

```
load("../assets/gerber_green_larimer.RData")
## turn turnout variable into a numeric
social$voted <- 1 * (social$voted == "Yes")
neigh.mean <- mean(social$voted[social$treatment == "Neighbors"])
neigh.mean
```

\#\# [1] 0.378

## Motivating Example

```
load("../assets/gerber_green_larimer.RData")
## turn turnout variable into a numeric
social$voted <- 1 * (social$voted == "Yes")
neigh.mean <- mean(social$voted[social$treatment == "Neighbors"])
neigh.mean
```

\#\# [1] 0.378
contr.mean <- mean(social\$voted[social\$treatment == "Civic Duty"])
contr.mean
\#\# [1] 0.315

## Motivating Example

```
load("../assets/gerber_green_larimer.RData")
## turn turnout variable into a numeric
social$voted <- 1 * (social$voted == "Yes")
neigh.mean <- mean(social$voted[social$treatment == "Neighbors"])
neigh.mean
```

\#\# [1] 0.378

```
contr.mean <- mean(social$voted[social$treatment == "Civic Duty"])
contr.mean
```

\#\# [1] 0.315
neigh.mean - contr.mean
\#\# [1] 0.0634

## Motivating Example

```
load("../assets/gerber_green_larimer.RData")
## turn turnout variable into a numeric
social$voted <- 1 * (social$voted == "Yes")
neigh.mean <- mean(social$voted[social$treatment == "Neighbors"])
neigh.mean
```

\#\# [1] 0.378

```
contr.mean <- mean(social$voted[social$treatment == "Civic Duty"])
contr.mean
```

\#\# [1] 0.315
neigh.mean - contr.mean
\#\# [1] 0.0634

- Is this a "real"? Is it big?


## Why study estimators?

- Goal 1: Inference


## Why study estimators?

- Goal 1: Inference
-What is our best guess about some quantity of interest?


## Why study estimators?

- Goal 1: Inference
-What is our best guess about some quantity of interest?
- What are a set of plausible values of the quantity of interest?


## Why study estimators?

- Goal 1: Inference
- What is our best guess about some quantity of interest?
- What are a set of plausible values of the quantity of interest?
- Goal 2: Compare estimators


## Why study estimators?

- Goal 1: Inference
- What is our best guess about some quantity of interest?
- What are a set of plausible values of the quantity of interest?
- Goal 2: Compare estimators
- In an experiment, use simple difference in sample means $(\bar{Y}-\bar{X})$ ?


## Why study estimators?

- Goal 1: Inference
- What is our best guess about some quantity of interest?
- What are a set of plausible values of the quantity of interest?
- Goal 2: Compare estimators
- In an experiment, use simple difference in sample means $(\bar{Y}-\bar{X})$ ?
- Or the post-stratification estimator, where we estimate the estimate the difference among two subsets of the data (male and female, for instance) and then take the weighted average of the two ( $\bar{Z}$ is the share of women):

$$
\left(\bar{Y}_{f}-\bar{X}_{f}\right) \bar{Z}+\left(\bar{Y}_{m}-\bar{X}_{m}\right)(1-\bar{Z})
$$

## Why study estimators?

- Goal 1: Inference
-What is our best guess about some quantity of interest?
- What are a set of plausible values of the quantity of interest?
- Goal 2: Compare estimators
- In an experiment, use simple difference in sample means $(\bar{Y}-\bar{X})$ ?
- Or the post-stratification estimator, where we estimate the estimate the difference among two subsets of the data (male and female, for instance) and then take the weighted average of the two ( $\bar{Z}$ is the share of women):

$$
\left(\bar{Y}_{f}-\bar{X}_{f}\right) \bar{Z}+\left(\bar{Y}_{m}-\bar{X}_{m}\right)(1-\bar{Z})
$$

- Which (if either) is better? How would we know?


## Samples from the population

- Model-based inferece: random vectors $X_{1}, \ldots, X_{n}$ are i.i.d. draws from c.d.f. F


## Samples from the population

- Model-based inferece: random vectors $X_{1}, \ldots, X_{n}$ are i.i.d. draws from c.d.f. F
- e.g.: $X_{i}=1$ if citizen $i$ votes, $X_{i}=0$ otherwise.


## Samples from the population

- Model-based inferece: random vectors $X_{1}, \ldots, X_{n}$ are i.i.d. draws from c.d.f. F
- e.g.: $X_{i}=1$ if citizen $i$ votes, $X_{i}=0$ otherwise.
- $n$ is the sample size


## Samples from the population

- Model-based inferece: random vectors $X_{1}, \ldots, X_{n}$ are i.i.d. draws from c.d.f. F
- e.g.: $X_{i}=1$ if citizen $i$ votes, $X_{i}=0$ otherwise.
- $n$ is the sample size
- i.i.d. can be justified through random sampling from an inifinite population.


## Samples from the population

- Model-based inferece: random vectors $X_{1}, \ldots, X_{n}$ are i.i.d. draws from c.d.f. $F$
- e.g.: $X_{i}=1$ if citizen $i$ votes, $X_{i}=0$ otherwise.
- $n$ is the sample size
- i.i.d. can be justified through random sampling from an inifinite population.
- $F$ is often called the population distribution or just population


## Samples from the population

- Model-based inferece: random vectors $X_{1}, \ldots, X_{n}$ are i.i.d. draws from c.d.f. $F$
- e.g.: $X_{i}=1$ if citizen $i$ votes, $X_{i}=0$ otherwise.
- $n$ is the sample size
- i.i.d. can be justified through random sampling from an inifinite population.
- $F$ is often called the population distribution or just population
- Model-based because we are assuming the probability model $F$


## Samples from the population

- Model-based inferece: random vectors $X_{1}, \ldots, X_{n}$ are i.i.d. draws from c.d.f. $F$
- e.g.: $X_{i}=1$ if citizen $i$ votes, $X_{i}=0$ otherwise.
- $n$ is the sample size
- i.i.d. can be justified through random sampling from an inifinite population.
- $F$ is often called the population distribution or just population
- Model-based because we are assuming the probability model $F$
- Two metaphors:


## Samples from the population

- Model-based inferece: random vectors $X_{1}, \ldots, X_{n}$ are i.i.d. draws from c.d.f. $F$
- e.g.: $X_{i}=1$ if citizen $i$ votes, $X_{i}=0$ otherwise.
- $n$ is the sample size
- i.i.d. can be justified through random sampling from an inifinite population.
- $F$ is often called the population distribution or just population
- Model-based because we are assuming the probability model $F$
- Two metaphors:
- Actual/potential population of size $N \gg n$ and we randomly sample $n$.


## Samples from the population

- Model-based inferece: random vectors $X_{1}, \ldots, X_{n}$ are i.i.d. draws from c.d.f. $F$
- e.g.: $X_{i}=1$ if citizen $i$ votes, $X_{i}=0$ otherwise.
- $n$ is the sample size
- i.i.d. can be justified through random sampling from an inifinite population.
- $F$ is often called the population distribution or just population
- Model-based because we are assuming the probability model $F$
- Two metaphors:
- Actual/potential population of size $N \gg n$ and we randomly sample $n$.
- $F$ represents the data generating process, we repeat $n$ times


## Samples from the population

- Model-based inferece: random vectors $X_{1}, \ldots, X_{n}$ are i.i.d. draws from c.d.f. $F$
- e.g.: $X_{i}=1$ if citizen $i$ votes, $X_{i}=0$ otherwise.
- $n$ is the sample size
- i.i.d. can be justified through random sampling from an inifinite population.
- $F$ is often called the population distribution or just population
- Model-based because we are assuming the probability model $F$
- Two metaphors:
- Actual/potential population of size $N \gg n$ and we randomly sample $n$.
- $F$ represents the data generating process, we repeat $n$ times
- Statistical inference or learning is using data to infer $F$.
- Goal of inference: learn about the features of the population.


## Point estimation

- Goal of inference: learn about the features of the population.
- Parameter: $\theta$ is any function of the population distribution $F$
- Goal of inference: learn about the features of the population.
- Parameter: $\theta$ is any function of the population distribution $F$
- Also called: quantities of interest, estimands.
- Goal of inference: learn about the features of the population.
- Parameter: $\theta$ is any function of the population distribution $F$
- Also called: quantities of interest, estimands.
- Examples of parameters:


## Point estimation

- Goal of inference: learn about the features of the population.
- Parameter: $\theta$ is any function of the population distribution $F$
- Also called: quantities of interest, estimands.
- Examples of parameters:
- $\mu=\mathbb{E}\left[X_{i}\right]$ : the mean (turnout rate in the population).


## Point estimation

- Goal of inference: learn about the features of the population.
- Parameter: $\theta$ is any function of the population distribution $F$
- Also called: quantities of interest, estimands.
- Examples of parameters:
- $\mu=\mathbb{E}\left[X_{i}\right]$ : the mean (turnout rate in the population).
- $\sigma^{2}=\mathbb{V}\left[X_{i}\right]$ : the variance.


## Point estimation

- Goal of inference: learn about the features of the population.
- Parameter: $\theta$ is any function of the population distribution $F$
- Also called: quantities of interest, estimands.
- Examples of parameters:
- $\mu=\mathbb{E}\left[X_{i}\right]$ : the mean (turnout rate in the population).
- $\sigma^{2}=\mathbb{V}\left[X_{i}\right]$ : the variance.
- $\mu_{y}-\mu_{x}=\mathbb{E}\left[Y_{i}\right]-\mathbb{E}\left[X_{i}\right]$ : the difference in mean turnout between two groups.


## Point estimation

- Goal of inference: learn about the features of the population.
- Parameter: $\theta$ is any function of the population distribution $F$
- Also called: quantities of interest, estimands.
- Examples of parameters:
- $\mu=\mathbb{E}\left[X_{i}\right]$ : the mean (turnout rate in the population).
- $\sigma^{2}=\mathbb{V}\left[X_{i}\right]$ : the variance.
- $\mu_{y}-\mu_{x}=\mathbb{E}\left[Y_{i}\right]-\mathbb{E}\left[X_{i}\right]$ : the difference in mean turnout between two groups.
- Point estimation: providing a single "best guess" about these parameters.


## Estimators

- A statistic is any function of the sample $\left\{X_{1}, \ldots, X_{n}\right\}$.


## Estimators

- A statistic is any function of the sample $\left\{X_{1}, \ldots, X_{n}\right\}$.
- Before we see the data, statistics are random and have distributions, etc.


## Estimators

- A statistic is any function of the sample $\left\{X_{1}, \ldots, X_{n}\right\}$.
- Before we see the data, statistics are random and have distributions, etc.
- After we see the data, statistic is realized and we see the specific value.


## Estimators

- A statistic is any function of the sample $\left\{X_{1}, \ldots, X_{n}\right\}$.
- Before we see the data, statistics are random and have distributions, etc.
- After we see the data, statistic is realized and we see the specific value.


## Definition

An estimator $\hat{\theta}_{n}$ for some parameter $\theta$, is a statistic intended as a guess about $\theta$.

## Estimators

- A statistic is any function of the sample $\left\{X_{1}, \ldots, X_{n}\right\}$.
- Before we see the data, statistics are random and have distributions, etc.
- After we see the data, statistic is realized and we see the specific value.


## Definition

An estimator $\hat{\theta}_{n}$ for some parameter $\theta$, is a statistic intended as a guess about $\theta$.

- $\hat{\theta}_{n}$ is a r.v. because it is a function of r.v.s.


## Estimators

- A statistic is any function of the sample $\left\{X_{1}, \ldots, X_{n}\right\}$.
- Before we see the data, statistics are random and have distributions, etc.
- After we see the data, statistic is realized and we see the specific value.


## Definition

An estimator $\hat{\theta}_{n}$ for some parameter $\theta$, is a statistic intended as a guess about $\theta$.

- $\hat{\theta}_{n}$ is a r.v. because it is a function of r.v.s.
- $\rightsquigarrow \hat{\theta}_{n}$ has a distribution.


## Estimators

- A statistic is any function of the sample $\left\{X_{1}, \ldots, X_{n}\right\}$.
- Before we see the data, statistics are random and have distributions, etc.
- After we see the data, statistic is realized and we see the specific value.


## Definition

An estimator $\hat{\theta}_{n}$ for some parameter $\theta$, is a statistic intended as a guess about $\theta$.

- $\hat{\theta}_{n}$ is a r.v. because it is a function of r.v.s.
- $\rightsquigarrow \hat{\theta}_{n}$ has a distribution.
- An estimate is one particular realization of the estimator


## Estimators

- A statistic is any function of the sample $\left\{X_{1}, \ldots, X_{n}\right\}$.
- Before we see the data, statistics are random and have distributions, etc.
- After we see the data, statistic is realized and we see the specific value.


## Definition

An estimator $\hat{\theta}_{n}$ for some parameter $\theta$, is a statistic intended as a guess about $\theta$.

- $\hat{\theta}_{n}$ is a r.v. because it is a function of r.v.s.
- $\rightsquigarrow \hat{\theta}_{n}$ has a distribution.
- An estimate is one particular realization of the estimator
- Why is the following statement wrong: "My estimate was the sample mean and my estimator was 0.38 "?


## Examples of Estimators

- For the population expectation, $\mathbb{E}\left[X_{i}\right]$, many possible estimators:


## Examples of Estimators

- For the population expectation, $\mathbb{E}\left[X_{i}\right]$, many possible estimators:
- $\hat{\theta}_{n}=\bar{X}_{n}$ the sample mean


## Examples of Estimators

- For the population expectation, $\mathbb{E}\left[X_{i}\right]$, many possible estimators:
- $\hat{\theta}_{n}=\bar{X}_{n}$ the sample mean
- $\hat{\theta}_{n}=X_{1}$ just use the first observation


## Examples of Estimators

- For the population expectation, $\mathbb{E}\left[X_{i}\right]$, many possible estimators:
- $\hat{\theta}_{n}=\bar{X}_{n}$ the sample mean
- $\hat{\theta}_{n}=X_{1}$ just use the first observation
- $\hat{\theta}_{n}=\max \left(X_{1}, \ldots, X_{n}\right)$


## Examples of Estimators

- For the population expectation, $\mathbb{E}\left[X_{i}\right]$, many possible estimators:
- $\hat{\theta}_{n}=\bar{X}_{n}$ the sample mean
- $\hat{\theta}_{n}=X_{1}$ just use the first observation
- $\hat{\theta}_{n}=\max \left(X_{1}, \ldots, X_{n}\right)$
- $\hat{\theta}_{n}=3$ always guess 3


## The three distributions

- Population Distribution: the data-generating process


## The three distributions

- Population Distribution: the data-generating process
- Bernoulli in the case of the social pressure/voter turnout example)


## The three distributions

- Population Distribution: the data-generating process
- Bernoulli in the case of the social pressure/voter turnout example)
- Empirical distribution: $X_{1}, \ldots, X_{n}$


## The three distributions

- Population Distribution: the data-generating process
- Bernoulli in the case of the social pressure/voter turnout example)
- Empirical distribution: $X_{1}, \ldots, X_{n}$
- series of 1 s and 0 s in the sample


## The three distributions

- Population Distribution: the data-generating process
- Bernoulli in the case of the social pressure/voter turnout example)
- Empirical distribution: $X_{1}, \ldots, X_{n}$
- series of 1 s and 0 s in the sample
- Sampling distribution: distribution of the estimator over repeated samples from the population distribution


## The three distributions

- Population Distribution: the data-generating process
- Bernoulli in the case of the social pressure/voter turnout example)
- Empirical distribution: $X_{1}, \ldots, X_{n}$
- series of 1 s and 0 s in the sample
- Sampling distribution: distribution of the estimator over repeated samples from the population distribution
- the 0.38 sample mean in the "Neighbors" group is one draw from this distribution


## Sampling distribution, in pictures


population

estimator distribution

## Sampling distribution, in pictures



## Sampling distribution, in pictures



## Sampling distribution, in pictures



## Sampling distribution, in pictures



## Sampling distribution

```
\#\# now we take the mean of one sample, which is
\#\# one draw from the **sampling distribution**
my.samp <- rbinom(n = 10, size = 1, prob = 0.4)
mean(my.samp)
```

\#\# [1] 0.4

## Sampling distribution

```
## now we take the mean of one sample, which is
## one draw from the **sampling distribution**
my.samp <- rbinom(n = 10, size = 1, prob = 0.4)
mean(my.samp)
```

\#\# [1] 0.4

```
## let's take another draw from the population dist
my.samp.2 <- rbinom(n = 10, size = 1, prob = 0.4)
```


## Sampling distribution

```
## now we take the mean of one sample, which is
## one draw from the **sampling distribution**
my.samp <- rbinom(n = 10, size = 1, prob = 0.4)
mean(my.samp)
```

\#\# [1] 0.4

```
## let's take another draw from the population dist
my.samp.2 <- rbinom(n = 10, size = 1, prob = 0.4)
```

\#\# Let's feed this sample to the sample mean estimator
\#\# to get another estimate, which is another draw from
\#\# the sampling distribution
mean(my.samp.2)
\#\# [1] 0.1

## Sampling distribution by simulation

- Let's generate 10,000 draws from the sampling distribution of the sample mean here when $n=100$.

```
nsims <- 10000
mean.holder <- rep(NA, times = nsims)
for (i in 1:nsims) {
    my.samp <- rbinom(n = 100, size = 1, prob = 0.4)
    mean.holder[i] <- mean(my.samp) ## sample mean
    first.holder[i] <- my.samp[1] ## first obs
}
```


## Sampling distribution versus population distribution



Question The sampling distribution refers to the distribution of $\theta$, true or false.

## Where do estimators come from?

- Parametric modeling: assume $X_{1}, \ldots, X_{n} \stackrel{\text { i.i.d. }}{\sim} F$ and specify what family $F$ is from.


## Where do estimators come from?

- Parametric modeling: assume $X_{1}, \ldots, X_{n} \stackrel{\text { i.i.d. }}{\sim} F$ and specify what family $F$ is from.
- Example: $F$ is $\operatorname{Pois}(\lambda)$.


## Where do estimators come from?

- Parametric modeling: assume $X_{1}, \ldots, X_{n} \stackrel{\text { i.i.d. }}{\sim} F$ and specify what family $F$ is from.
- Example: $F$ is $\operatorname{Pois}(\lambda)$.
- Construct estimator $\hat{\lambda}$ using maximum likelihood


## Where do estimators come from?

- Parametric modeling: assume $X_{1}, \ldots, X_{n} \stackrel{\text { i.i.d. }}{\sim} F$ and specify what family $F$ is from.
- Example: $F$ is $\operatorname{Pois}(\lambda)$.
- Construct estimator $\hat{\lambda}$ using maximum likelihood
- Downside: inferences are model dependent


## Where do estimators come from?

- Parametric modeling: assume $X_{1}, \ldots, X_{n} \stackrel{\text { i.i.d. }}{\sim} F$ and specify what family $F$ is from.
- Example: $F$ is $\operatorname{Pois}(\lambda)$.
- Construct estimator $\hat{\lambda}$ using maximum likelihood
- Downside: inferences are model dependent
- Nonparametric inference: make minimal assumptions on $F$.


## Where do estimators come from?

- Parametric modeling: assume $X_{1}, \ldots, X_{n} \stackrel{\text { i.i.d. }}{\sim} F$ and specify what family $F$ is from.
- Example: $F$ is $\operatorname{Pois}(\lambda)$.
- Construct estimator $\hat{\lambda}$ using maximum likelihood
- Downside: inferences are model dependent
- Nonparametric inference: make minimal assumptions on $F$.
- Plug-in/analogy principle: replace $F$ with the empirical distribution.


## Where do estimators come from?

- Parametric modeling: assume $X_{1}, \ldots, X_{n} \stackrel{\text { i.i.d. }}{\sim} F$ and specify what family $F$ is from.
- Example: $F$ is $\operatorname{Pois}(\lambda)$.
- Construct estimator $\hat{\lambda}$ using maximum likelihood
- Downside: inferences are model dependent
- Nonparametric inference: make minimal assumptions on $F$.
- Plug-in/analogy principle: replace $F$ with the empirical distribution.
- Empirical distribution: probability $1 / n$ at each observed value of $X_{i}$ :

$$
\widehat{F}_{n}(x)=\frac{\sum_{i=1}^{n} \mathbb{\square}\left(X_{i} \leq x\right)}{n}
$$

## Where do estimators come from?

- Parametric modeling: assume $X_{1}, \ldots, X_{n} \stackrel{\text { i.i.d. }}{\sim} F$ and specify what family $F$ is from.
- Example: $F$ is $\operatorname{Pois}(\lambda)$.
- Construct estimator $\hat{\lambda}$ using maximum likelihood
- Downside: inferences are model dependent
- Nonparametric inference: make minimal assumptions on $F$.
- Plug-in/analogy principle: replace $F$ with the empirical distribution.
- Empirical distribution: probability $1 / n$ at each observed value of $X_{i}$ :

$$
\widehat{F}_{n}(x)=\frac{\sum_{i=1}^{n} \square\left(X_{i} \leq x\right)}{n}
$$

- $\rightsquigarrow$ if $\theta=\mathbb{E}[g(X)]$ replace $\mathbb{E}$ sample means: $\hat{\theta}=\frac{1}{n} \sum_{i=1}^{n} g\left(X_{i}\right)$


## Plug-in estimators, examples

- Expectation:

$$
\mu=\mathbb{E}\left[X_{i}\right] \rightsquigarrow \hat{\mu}=\frac{1}{n} \sum_{i=1}^{n} X_{i}=\bar{X}_{n}
$$

## Plug-in estimators, examples

- Expectation:

$$
\mu=\mathbb{E}\left[X_{i}\right] \rightsquigarrow \hat{\mu}=\frac{1}{n} \sum_{i=1}^{n} X_{i}=\bar{X}_{n}
$$

- Variance:

$$
\sigma^{2}=\mathbb{E}\left[\left(X_{i}-\mathbb{E}\left[X_{i}\right]\right)^{2}\right] \rightsquigarrow \widehat{\sigma}^{2}=\frac{1}{n} \sum_{i=1}^{n}\left(X_{i}-\bar{X}_{n}\right)^{2}
$$

## Plug-in estimators, examples

- Expectation:

$$
\mu=\mathbb{E}\left[X_{i}\right] \rightsquigarrow \hat{\mu}=\frac{1}{n} \sum_{i=1}^{n} X_{i}=\bar{X}_{n}
$$

- Variance:

$$
\sigma^{2}=\mathbb{E}\left[\left(X_{i}-\mathbb{E}\left[X_{i}\right]\right)^{2}\right] \rightsquigarrow \widehat{\sigma}^{2}=\frac{1}{n} \sum_{i=1}^{n}\left(X_{i}-\bar{X}_{n}\right)^{2}
$$

- Covariance:

$$
\sigma_{x y}=\operatorname{Cov}\left[X_{i}, Y_{i}\right]=\mathbb{E}\left[\left(X_{i}-\mathbb{E}\left[X_{i}\right]\right)\left(Y_{i}-\mathbb{E}\left[Y_{i}\right]\right)\right] \rightsquigarrow \widehat{\sigma}_{x y}=\frac{1}{n} \sum_{i=1}^{n}\left(X_{i}-\bar{X}\right)\left(Y_{i}-\bar{Y}\right)
$$

2/ Finite-Sample
Properties of Estimators

## Properties of estimators

- We only get one draw from the sampling distribution, $\hat{\theta}_{n}$.


## Properties of estimators

- We only get one draw from the sampling distribution, $\hat{\theta}_{n}$.
- Want to use estimators whose distribution is "close" to the true value.


## Properties of estimators

- We only get one draw from the sampling distribution, $\hat{\theta}_{n}$.
- Want to use estimators whose distribution is "close" to the true value.
- There are two ways we evaluate estimators:


## Properties of estimators

- We only get one draw from the sampling distribution, $\hat{\theta}_{n}$.
- Want to use estimators whose distribution is "close" to the true value.
- There are two ways we evaluate estimators:
- Finite sample: the properties of its sampling distribution for a fixed sample size $n$.


## Properties of estimators

- We only get one draw from the sampling distribution, $\hat{\theta}_{n}$.
- Want to use estimators whose distribution is "close" to the true value.
- There are two ways we evaluate estimators:
- Finite sample: the properties of its sampling distribution for a fixed sample size $n$.
- Large sample: the properties of the sampling distribution as we let $n \rightarrow \infty$.
- The bias of estimator $\hat{\theta}$ for parameter $\theta$ is

$$
\operatorname{bias}[\hat{\theta}]=\mathbb{E}[\hat{\theta}]-\theta .
$$

- The bias of estimator $\hat{\theta}$ for parameter $\theta$ is

$$
\operatorname{bias}[\hat{\theta}]=\mathbb{E}[\hat{\theta}]-\theta \text {. }
$$

- An estimator is unbiased if $\operatorname{bias}[\hat{\theta}]=0$.
- The bias of estimator $\hat{\theta}$ for parameter $\theta$ is

$$
\operatorname{bias}[\hat{\theta}]=\mathbb{E}[\hat{\theta}]-\theta .
$$

- An estimator is unbiased if bias $[\hat{\theta}]=0$.
- Sample mean of i.i.d. $X_{1}, \ldots, X_{n}$ with $\mathbb{E}\left[X_{i}\right]=\mu$

$$
\mathbb{E}\left[\bar{X}_{n}\right]=\frac{1}{n} \sum_{i=1}^{n} \mathbb{E}\left[X_{i}\right]=\frac{1}{n} \sum_{i=1}^{n} \mu=\mu
$$

- The bias of estimator $\hat{\theta}$ for parameter $\theta$ is

$$
\operatorname{bias}[\hat{\theta}]=\mathbb{E}[\hat{\theta}]-\theta .
$$

- An estimator is unbiased if bias $[\hat{\theta}]=0$.
- Sample mean of i.i.d. $X_{1}, \ldots, X_{n}$ with $\mathbb{E}\left[X_{i}\right]=\mu$

$$
\mathbb{E}\left[\bar{X}_{n}\right]=\frac{1}{n} \sum_{i=1}^{n} \mathbb{E}\left[X_{i}\right]=\frac{1}{n} \sum_{i=1}^{n} \mu=\mu
$$

- Thus, $\bar{X}_{n}$ is unbiased for $\mu$ if $\mathbb{E}[|X|]<\infty$
- The bias of estimator $\hat{\theta}$ for parameter $\theta$ is

$$
\operatorname{bias}[\hat{\theta}]=\mathbb{E}[\hat{\theta}]-\theta .
$$

- An estimator is unbiased if bias $[\hat{\theta}]=0$.
- Sample mean of i.i.d. $X_{1}, \ldots, X_{n}$ with $\mathbb{E}\left[X_{i}\right]=\mu$

$$
\mathbb{E}\left[\bar{X}_{n}\right]=\frac{1}{n} \sum_{i=1}^{n} \mathbb{E}\left[X_{i}\right]=\frac{1}{n} \sum_{i=1}^{n} \mu=\mu
$$

- Thus, $\bar{X}_{n}$ is unbiased for $\mu$ if $\mathbb{E}[|X|]<\infty$
- What about a weighted average?
- The bias of estimator $\hat{\theta}$ for parameter $\theta$ is

$$
\operatorname{bias}[\hat{\theta}]=\mathbb{E}[\hat{\theta}]-\theta .
$$

- An estimator is unbiased if bias $[\hat{\theta}]=0$.
- Sample mean of i.i.d. $X_{1}, \ldots, X_{n}$ with $\mathbb{E}\left[X_{i}\right]=\mu$

$$
\mathbb{E}\left[\bar{X}_{n}\right]=\frac{1}{n} \sum_{i=1}^{n} \mathbb{E}\left[X_{i}\right]=\frac{1}{n} \sum_{i=1}^{n} \mu=\mu
$$

- Thus, $\bar{X}_{n}$ is unbiased for $\mu$ if $\mathbb{E}[|X|]<\infty$
- What about a weighted average?
- Unbiasedness is preserved under linear transformations.


## Estimation variance

- Sampling variance: the variance of an estimator $\mathbb{V}[\hat{\theta}]$.


## Estimation variance

- Sampling variance: the variance of an estimator $\mathbb{\vee}[\hat{\theta}]$.
- Measure of how spread the estimator it is around its mean.


## Estimation variance

- Sampling variance: the variance of an estimator $\vee[\hat{\theta}]$.
- Measure of how spread the estimator it is around its mean.
- Sampling variance of the sample mean:

$$
\mathbb{V}\left[\bar{X}_{n}\right]=\frac{1}{n^{2}} \sum_{i=1}^{n} \mathbb{V}\left[X_{i}\right]=\frac{1}{n^{2}} \sum_{i=1}^{n} \sigma^{2}=\frac{\sigma^{2}}{n}
$$

## Estimation variance

- Sampling variance: the variance of an estimator $\mathbb{V}[\hat{\theta}]$.
- Measure of how spread the estimator it is around its mean.
- Sampling variance of the sample mean:

$$
\mathbb{V}\left[\bar{X}_{n}\right]=\frac{1}{n^{2}} \sum_{i=1}^{n} \mathbb{V}\left[X_{i}\right]=\frac{1}{n^{2}} \sum_{i=1}^{n} \sigma^{2}=\frac{\sigma^{2}}{n}
$$

- Standard error: standard deviation of the estimator se $(\hat{\theta})=\sqrt{\vee[\hat{\theta}]}$


## Estimation variance

- Sampling variance: the variance of an estimator $\mathbb{\vee}[\hat{\theta}]$.
- Measure of how spread the estimator it is around its mean.
- Sampling variance of the sample mean:

$$
\mathbb{V}\left[\bar{X}_{n}\right]=\frac{1}{n^{2}} \sum_{i=1}^{n} \mathbb{V}\left[X_{i}\right]=\frac{1}{n^{2}} \sum_{i=1}^{n} \sigma^{2}=\frac{\sigma^{2}}{n}
$$

- Standard error: standard deviation of the estimator se $(\hat{\theta})=\sqrt{\vee[\hat{\theta}]}$
- Like all SDs, nice that it's on the same scale.


## Estimation variance

- Sampling variance: the variance of an estimator $\vee[\hat{\theta}]$.
- Measure of how spread the estimator it is around its mean.
- Sampling variance of the sample mean:

$$
\mathbb{V}\left[\bar{X}_{n}\right]=\frac{1}{n^{2}} \sum_{i=1}^{n} \mathbb{V}\left[X_{i}\right]=\frac{1}{n^{2}} \sum_{i=1}^{n} \sigma^{2}=\frac{\sigma^{2}}{n}
$$

- Standard error: standard deviation of the estimator $\operatorname{se}(\hat{\theta})=\sqrt{\vee[\hat{\theta}]}$
- Like all SDs, nice that it's on the same scale.
- Standard error of the sample mean: $\sigma / \sqrt{n}$


## Mean squared error

- Mean squared error or MSE is

$$
\mathrm{MSE}=\mathbb{E}\left[\left(\hat{\theta}_{n}-\theta\right)^{2}\right]
$$

## Mean squared error

- Mean squared error or MSE is

$$
\mathrm{MSE}=\mathbb{E}\left[\left(\hat{\theta}_{n}-\theta\right)^{2}\right]
$$

- The MSE assesses the quality of an estimator.


## Mean squared error

- Mean squared error or MSE is

$$
\mathrm{MSE}=\mathbb{E}\left[\left(\hat{\theta}_{n}-\theta\right)^{2}\right]
$$

- The MSE assesses the quality of an estimator.
- How big are (squared) deviations from the true parameter?


## Mean squared error

- Mean squared error or MSE is

$$
\text { MSE }=\mathbb{E}\left[\left(\hat{\theta}_{n}-\theta\right)^{2}\right]
$$

- The MSE assesses the quality of an estimator.
- How big are (squared) deviations from the true parameter?
- Ideally, this would be as low as possible!


## Mean squared error

- Mean squared error or MSE is

$$
\mathrm{MSE}=\mathbb{E}\left[\left(\hat{\theta}_{n}-\theta\right)^{2}\right]
$$

- The MSE assesses the quality of an estimator.
- How big are (squared) deviations from the true parameter?
- Ideally, this would be as low as possible!
- Useful decomposition result:

$$
\mathrm{MSE}=\operatorname{bias}\left[\hat{\theta}_{n}\right]^{2}+V\left[\hat{\theta}_{n}\right]
$$

## Mean squared error

- Mean squared error or MSE is

$$
\text { MSE }=\mathbb{E}\left[\left(\hat{\theta}_{n}-\theta\right)^{2}\right]
$$

- The MSE assesses the quality of an estimator.
- How big are (squared) deviations from the true parameter?
- Ideally, this would be as low as possible!
- Useful decomposition result:

$$
\mathrm{MSE}=\operatorname{bias}\left[\hat{\theta}_{n}\right]^{2}+\mathbb{V}\left[\hat{\theta}_{n}\right]
$$

- $\rightsquigarrow$ for unbiased estimators, MSE is the sampling variance.


## Mean squared error

- Mean squared error or MSE is

$$
\text { MSE }=\mathbb{E}\left[\left(\hat{\theta}_{n}-\theta\right)^{2}\right]
$$

- The MSE assesses the quality of an estimator.
- How big are (squared) deviations from the true parameter?
- Ideally, this would be as low as possible!
- Useful decomposition result:

$$
\mathrm{MSE}=\operatorname{bias}\left[\hat{\theta}_{n}\right]^{2}+\mathbb{V}\left[\hat{\theta}_{n}\right]
$$

- $\rightsquigarrow$ for unbiased estimators, MSE is the sampling variance.
- Might accept some bias for large reductions in variance for lower overall MSE.

3/ Design-based inference

## Survey sampling

- Up to now: focus on model-based inference.


## Survey sampling

- Up to now: focus on model-based inference.
- $X_{1}, \ldots, X_{n}$ are iid draws from an infinite population modeled by $\operatorname{cdf} F$


## Survey sampling

- Up to now: focus on model-based inference.
- $X_{1}, \ldots, X_{n}$ are iid draws from an infinite population modeled by $c d f F$
- Alternative: a large, but finite sample of size $N$ indexed $i=1, \ldots, N$.


## Survey sampling

- Up to now: focus on model-based inference.
- $X_{1}, \ldots, X_{n}$ are iid draws from an infinite population modeled by $c d f F$
- Alternative: a large, but finite sample of size $N$ indexed $i=1, \ldots, N$.
- Population characteristics: $x_{1}, x_{2}, \ldots, x_{N}$ (list of fixed numbers)


## Survey sampling

- Up to now: focus on model-based inference.
- $X_{1}, \ldots, X_{n}$ are iid draws from an infinite population modeled by $c d f F$
- Alternative: a large, but finite sample of size $N$ indexed $i=1, \ldots, N$.
- Population characteristics: $x_{1}, x_{2}, \ldots, x_{N}$ (list of fixed numbers)
- We'll think of the population and everything about it as fixed


## Survey sampling

- Up to now: focus on model-based inference.
- $X_{1}, \ldots, X_{n}$ are iid draws from an infinite population modeled by cdf $F$
- Alternative: a large, but finite sample of size $N$ indexed $i=1, \ldots, N$.
- Population characteristics: $x_{1}, x_{2}, \ldots, x_{N}$ (list of fixed numbers)
- We'll think of the population and everything about it as fixed
- Assumption: simple random sample (eg, with replacement) of size $n$ from this population


## Survey sampling

- Up to now: focus on model-based inference.
- $X_{1}, \ldots, X_{n}$ are iid draws from an infinite population modeled by cdf $F$
- Alternative: a large, but finite sample of size $N$ indexed $i=1, \ldots, N$.
- Population characteristics: $x_{1}, x_{2}, \ldots, x_{N}$ (list of fixed numbers)
- We'll think of the population and everything about it as fixed
- Assumption: simple random sample (eg, with replacement) of size $n$ from this population
- Number of possible samples: $\binom{N}{n}$


## Survey sampling

- Up to now: focus on model-based inference.
- $X_{1}, \ldots, X_{n}$ are iid draws from an infinite population modeled by $\operatorname{cdf} F$
- Alternative: a large, but finite sample of size $N$ indexed $i=1, \ldots, N$.
- Population characteristics: $x_{1}, x_{2}, \ldots, x_{N}$ (list of fixed numbers)
- We'll think of the population and everything about it as fixed
- Assumption: simple random sample (eg, with replacement) of size $n$ from this population
- Number of possible samples: $\binom{N}{n}$
- Sampling inclusion indicators: $I_{1}, I_{2}, \ldots, I_{N}$


## Survey sampling

- Up to now: focus on model-based inference.
- $X_{1}, \ldots, X_{n}$ are iid draws from an infinite population modeled by $c d f F$
- Alternative: a large, but finite sample of size $N$ indexed $i=1, \ldots, N$.
- Population characteristics: $x_{1}, x_{2}, \ldots, x_{N}$ (list of fixed numbers)
- We'll think of the population and everything about it as fixed
- Assumption: simple random sample (eg, with replacement) of size $n$ from this population
- Number of possible samples: $\binom{N}{n}$
- Sampling inclusion indicators: $I_{1}, I_{2}, \ldots, I_{N}$
- These are random because of the random sampling (uppercase!)


## Survey sampling

- Up to now: focus on model-based inference.
- $X_{1}, \ldots, X_{n}$ are iid draws from an infinite population modeled by $\operatorname{cdf} F$
- Alternative: a large, but finite sample of size $N$ indexed $i=1, \ldots, N$.
- Population characteristics: $x_{1}, x_{2}, \ldots, x_{N}$ (list of fixed numbers)
- We'll think of the population and everything about it as fixed
- Assumption: simple random sample (eg, with replacement) of size $n$ from this population
- Number of possible samples: $\binom{N}{n}$
- Sampling inclusion indicators: $I_{1}, I_{2}, \ldots, I_{N}$
- These are random because of the random sampling (uppercase!)
- Total sample size is fixed: $\sum_{i=1}^{N} I_{i}=n$


## Survey sampling

- Up to now: focus on model-based inference.
- $X_{1}, \ldots, X_{n}$ are iid draws from an infinite population modeled by $\operatorname{cdf} F$
- Alternative: a large, but finite sample of size $N$ indexed $i=1, \ldots, N$.
- Population characteristics: $x_{1}, x_{2}, \ldots, x_{N}$ (list of fixed numbers)
- We'll think of the population and everything about it as fixed
- Assumption: simple random sample (eg, with replacement) of size $n$ from this population
- Number of possible samples: $\binom{N}{n}$
- Sampling inclusion indicators: $I_{1}, I_{2}, \ldots, I_{N}$
- These are random because of the random sampling (uppercase!)
- Total sample size is fixed: $\sum_{i=1}^{N} l_{i}=n$
- Inclusion probabilities: $\pi=\mathbb{P}\left(I_{i}=1\right)=n / N$


## Survey sampling

- Up to now: focus on model-based inference.
- $X_{1}, \ldots, X_{n}$ are iid draws from an infinite population modeled by $\operatorname{cdf} F$
- Alternative: a large, but finite sample of size $N$ indexed $i=1, \ldots, N$.
- Population characteristics: $x_{1}, x_{2}, \ldots, x_{N}$ (list of fixed numbers)
- We'll think of the population and everything about it as fixed
- Assumption: simple random sample (eg, with replacement) of size $n$ from this population
- Number of possible samples: $\binom{N}{n}$
- Sampling inclusion indicators: $I_{1}, I_{2}, \ldots, I_{N}$
- These are random because of the random sampling (uppercase!)
- Total sample size is fixed: $\sum_{i=1}^{N} I_{i}=n$
- Inclusion probabilities: $\pi=\mathbb{P}\left(I_{i}=1\right)=n / N$
- Different sampling designs lead to different inclusion probabilities and difference inferences.


## Estimands and estimators

- Estimand: population mean $\bar{x}=\frac{1}{N} \sum_{i=1}^{N} x_{i}$


## Estimands and estimators

- Estimand: population mean $\bar{x}=\frac{1}{N} \sum_{i=1}^{N} x_{i}$
- Fixed quantity because the population is fixed and finite.


## Estimands and estimators

- Estimand: population mean $\bar{x}=\frac{1}{N} \sum_{i=1}^{N} x_{i}$
- Fixed quantity because the population is fixed and finite.
- But we don't observe all $x_{i}$, so we cannot calculate it.


## Estimands and estimators

- Estimand: population mean $\bar{x}=\frac{1}{N} \sum_{i=1}^{N} x_{i}$
- Fixed quantity because the population is fixed and finite.
- But we don't observe all $x_{i}$, so we cannot calculate it.
- Estimator: sample mean $\bar{X}_{n}=\frac{1}{n} \sum_{i=1}^{N} l_{i} x_{i}$


## Estimands and estimators

- Estimand: population mean $\bar{x}=\frac{1}{N} \sum_{i=1}^{N} x_{i}$
- Fixed quantity because the population is fixed and finite.
- But we don't observe all $x_{i}$, so we cannot calculate it.
- Estimator: sample mean $\bar{X}_{n}=\frac{1}{n} \sum_{i=1}^{N} l_{i} x_{i}$
- This estimator is random because the sample is random.


## Estimands and estimators

- Estimand: population mean $\bar{x}=\frac{1}{N} \sum_{i=1}^{N} x_{i}$
- Fixed quantity because the population is fixed and finite.
- But we don't observe all $x_{i}$, so we cannot calculate it.
- Estimator: sample mean $\bar{X}_{n}=\frac{1}{n} \sum_{i=1}^{N} l_{i} x_{i}$
- This estimator is random because the sample is random.
- Design-based inference: randomness comes from sampling alone and depends on sampling design.


## Estimands and estimators

- Estimand: population mean $\bar{x}=\frac{1}{N} \sum_{i=1}^{N} x_{i}$
- Fixed quantity because the population is fixed and finite.
- But we don't observe all $x_{i}$, so we cannot calculate it.
- Estimator: sample mean $\bar{X}_{n}=\frac{1}{n} \sum_{i=1}^{N} l_{i} x_{i}$
- This estimator is random because the sample is random.
- Design-based inference: randomness comes from sampling alone and depends on sampling design.
- Unbiasedness proof is illustrative:

$$
\mathbb{E}\left[\bar{X}_{n}\right]=\mathbb{E}\left[\frac{1}{n} \sum_{i=1}^{N} I_{i} x_{i}\right]=\frac{1}{n} \sum_{i=1}^{N} \mathbb{E}\left[l_{i}\right] x_{i}=\frac{1}{n} \sum_{i=1}^{N} \frac{n}{N} x_{i}=\frac{1}{N} \sum_{i=1}^{N} x_{i}=\bar{x}
$$

## Estimands and estimators

- Estimand: population mean $\bar{x}=\frac{1}{N} \sum_{i=1}^{N} x_{i}$
- Fixed quantity because the population is fixed and finite.
- But we don't observe all $x_{i}$, so we cannot calculate it.
- Estimator: sample mean $\bar{X}_{n}=\frac{1}{n} \sum_{i=1}^{N} l_{i} x_{i}$
- This estimator is random because the sample is random.
- Design-based inference: randomness comes from sampling alone and depends on sampling design.
- Unbiasedness proof is illustrative:

$$
\mathbb{E}\left[\bar{X}_{n}\right]=\mathbb{E}\left[\frac{1}{n} \sum_{i=1}^{N} I_{i} x_{i}\right]=\frac{1}{n} \sum_{i=1}^{N} \mathbb{E}\left[l_{i}\right] x_{i}=\frac{1}{n} \sum_{i=1}^{N} \frac{n}{N} x_{i}=\frac{1}{N} \sum_{i=1}^{N} x_{i}=\bar{x}
$$

- Remember: unbiased across repeated samples from the sampling design.


## Variance of the sample mean

- Variance of $\bar{X}_{n}$ across repeated samples:

$$
\mathbb{V}\left[\bar{X}_{n}\right]=\underbrace{\left(1-\frac{n}{N}\right)}_{\text {finite pop. correction }} \frac{s^{2}}{n}
$$

## Variance of the sample mean

- Variance of $\bar{X}_{n}$ across repeated samples:

$$
V\left[\bar{X}_{n}\right]=\underbrace{\left(1-\frac{n}{N}\right)}_{\text {finite pop. correction }} \frac{s^{2}}{n}
$$

- $s^{2}$ is the population variance of $x_{i}$ (a fixed quantity!!):

$$
s^{2}=\frac{1}{N-1} \sum_{i=1}^{N}\left(x_{i}-\bar{x}\right)^{2}
$$

## Variance of the sample mean

- Variance of $\bar{X}_{n}$ across repeated samples:

$$
V\left[\bar{X}_{n}\right]=\underbrace{\left(1-\frac{n}{N}\right)}_{\text {finite pop. correction }} \frac{s^{2}}{n}
$$

- $s^{2}$ is the population variance of $x_{i}$ (a fixed quantity!!):

$$
s^{2}=\frac{1}{N-1} \sum_{i=1}^{N}\left(x_{i}-\bar{x}\right)^{2}
$$

- We can still apply the plug-in principle and use the sample variance $S^{2}$

$$
\hat{\mathbb{V}}\left[\bar{X}_{n}\right]=\left(1-\frac{n}{N}\right) \frac{S^{2}}{n} \quad S^{2}=\frac{1}{n-1} \sum_{i=1}^{N} I_{i}\left(x_{i}-\bar{X}_{n}\right)^{2}
$$

## Variance of the sample mean

- Variance of $\bar{X}_{n}$ across repeated samples:

$$
V\left[\bar{X}_{n}\right]=\underbrace{\left(1-\frac{n}{N}\right)}_{\text {finite pop. correction }} \frac{s^{2}}{n}
$$

- $s^{2}$ is the population variance of $x_{i}$ (a fixed quantity!!):

$$
s^{2}=\frac{1}{N-1} \sum_{i=1}^{N}\left(x_{i}-\bar{x}\right)^{2}
$$

- We can still apply the plug-in principle and use the sample variance $S^{2}$

$$
\hat{\mathbb{V}}\left[\bar{X}_{n}\right]=\left(1-\frac{n}{N}\right) \frac{S^{2}}{n} \quad S^{2}=\frac{1}{n-1} \sum_{i=1}^{N} I_{i}\left(x_{i}-\bar{X}_{n}\right)^{2}
$$

- We can show that this is unbiased so that $\mathbb{E}\left[\hat{\mathbb{V}}\left[\bar{X}_{n}\right]\right]=\mathbb{V}\left[\bar{X}_{n}\right]$


## Inverse probability weighting

- More often, we have unequal sampling probabilities: $\pi_{i}=\mathbb{P}\left(I_{i}=1\right)$ for each $i$


## Inverse probability weighting

- More often, we have unequal sampling probabilities: $\pi_{i}=\mathbb{P}\left(I_{i}=1\right)$ for each $i$
- Typically to oversample groups that are difficult to reach


## Inverse probability weighting

- More often, we have unequal sampling probabilities: $\pi_{i}=\mathbb{P}\left(I_{i}=1\right)$ for each $i$
- Typically to oversample groups that are difficult to reach
- Or to ensure sufficient sample sizes for smaller minority groups


## Inverse probability weighting

- More often, we have unequal sampling probabilities: $\pi_{i}=\mathbb{P}\left(I_{i}=1\right)$ for each $i$
- Typically to oversample groups that are difficult to reach
- Or to ensure sufficient sample sizes for smaller minority groups
- Horvitz-Thompson estimator:

$$
\widetilde{X}_{H T}=\frac{1}{N} \sum_{i=1}^{N} \frac{I_{i} x_{i}}{\pi_{i}}
$$

## Inverse probability weighting

- More often, we have unequal sampling probabilities: $\pi_{i}=\mathbb{P}\left(I_{i}=1\right)$ for each $i$
- Typically to oversample groups that are difficult to reach
- Or to ensure sufficient sample sizes for smaller minority groups
- Horvitz-Thompson estimator:

$$
\widetilde{X}_{H T}=\frac{1}{N} \sum_{i=1}^{N} \frac{l_{i} x_{i}}{\pi_{i}}
$$

- The HT estimator is unbiased: $\mathbb{E}\left[\widetilde{X}_{H T}\right]=\bar{x}$


## Inverse probability weighting

- More often, we have unequal sampling probabilities: $\pi_{i}=\mathbb{P}\left(I_{i}=1\right)$ for each $i$
- Typically to oversample groups that are difficult to reach
- Or to ensure sufficient sample sizes for smaller minority groups
- Horvitz-Thompson estimator:

$$
\widetilde{X}_{H T}=\frac{1}{N} \sum_{i=1}^{N} \frac{l_{i} x_{i}}{\pi_{i}}
$$

- The HT estimator is unbiased: $\mathbb{E}\left[\widetilde{X}_{H T}\right]=\bar{x}$
- But be very unstable and high variance if a low $\pi_{i}$ actually gets sampled


## Inverse probability weighting

- More often, we have unequal sampling probabilities: $\pi_{i}=\mathbb{P}\left(I_{i}=1\right)$ for each $i$
- Typically to oversample groups that are difficult to reach
- Or to ensure sufficient sample sizes for smaller minority groups
- Horvitz-Thompson estimator:

$$
\widetilde{X}_{H T}=\frac{1}{N} \sum_{i=1}^{N} \frac{l_{i} x_{i}}{\pi_{i}}
$$

- The HT estimator is unbiased: $\mathbb{E}\left[\widetilde{X}_{H T}\right]=\bar{x}$
- But be very unstable and high variance if a low $\pi_{i}$ actually gets sampled
- Alternative: Hajek estimator (also known as the IPW estimator)

$$
\widetilde{X}_{i p w}=\frac{\sum_{i=1}^{N} l_{i} x_{i} / \pi_{i}}{\sum_{i=1}^{N} l_{i} / \pi_{i}}
$$

## Inverse probability weighting

- More often, we have unequal sampling probabilities: $\pi_{i}=\mathbb{P}\left(I_{i}=1\right)$ for each $i$
- Typically to oversample groups that are difficult to reach
- Or to ensure sufficient sample sizes for smaller minority groups
- Horvitz-Thompson estimator:

$$
\widetilde{X}_{H T}=\frac{1}{N} \sum_{i=1}^{N} \frac{l_{i} x_{i}}{\pi_{i}}
$$

- The HT estimator is unbiased: $\mathbb{E}\left[\widetilde{X}_{H T}\right]=\bar{x}$
- But be very unstable and high variance if a low $\pi_{i}$ actually gets sampled
- Alternative: Hajek estimator (also known as the IPW estimator)

$$
\widetilde{X}_{i p w}=\frac{\sum_{i=1}^{N} l_{i} x_{i} / \pi_{i}}{\sum_{i=1}^{N} I_{i} / \pi_{i}}
$$

- Normalizes by the sum of weights.

