10. Hypothesis Testing

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Gov 2002 (Harvard)

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- We'll draw on our probability knowledge from earlier in the term!

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 - This is our data. What can we learn from it?
 - There is uncertainty: she could have guessed randomly.

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- \rightsquigarrow the guessing at random hypothesis might be implausible.

Social pressure effect

Dear Registered Voter:

WHAT IF YOUR NEIGHBORS KNEW WHETHER YOU VOTED?

Why do so many people fail to vote? We've been talking about the problem for years, but it only seems to get worse. This year, we're taking a new approach. We're sending this mailing to you and your neighbors to publicize who does and does not vote.

The chart shows the names of some of your neighbors, showing which have voted in the past. After the August 8 election, we intend to mail an updated chart. You and your neighbors will all know who voted and who did not.

DO YOUR CIVIC DUTY - VOTE!

MAPLE DR	Aug 04	Nov 04	Aug 06
9995 JOSEPH JAMES SMITH	Voted	Voted	
9995 JENNIFER KAY SMITH		Voted	
9997 RICHARD B JACKSON		Voted	
9999 KATHY MARIE JACKSON		Voted	

TABLE 2. Effects of Four Mail Treatments on Voter Turnout in the August 2006 Primary Election

	Experimental Group					
	Control	Civic Duty	Hawthorne	Self	Neighbors	
Percentage Voting	29.7%	31.5%	32.2%	34.5%	37.8%	
N of Individuals	191,243	38,218	38,204	38,218	38,201	

```
load("../assets/gerber_green_larimer.RData")
social$voted <- 1 * (social$voted == "Yes")
neigh.mean <- mean(social$voted[social$treatment == "Neighbors"])
contr.mean <- mean(social$voted[social$treatment == "Civic Duty"])
neigh.mean - contr.mean</pre>
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- Treatment effect of 6.3 percentage points.
- But the estimator varies from sample to sample by random chance.
- Could it be this big by random chance if there was no effect at all?

• **Treated group** $Y_1, Y_2, ..., Y_{n_y}$ i.i.d. with population mean μ_y and population variance σ_y^2

- Treated group Y_1, Y_2, \dots, Y_{n_y} i.i.d. with population mean μ_y and population variance σ_y^2
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- Estimator: sample difference in means: $\hat{\tau}_n = \overline{Y}_{n_v} \overline{X}_{n_x}$
- We estimate the standard error of $\hat{\tau}_n$ with:

$$\widehat{\mathsf{se}}[\widehat{\tau}_n] = \sqrt{\frac{s_y^2}{n_y} + \frac{s_x^2}{n_x}}$$

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 - · Are traits of treatment and control groups different?

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- 5. Reject if T_n in rejection region, fail to reject otherwise

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 - Two-sided much more common, one-sided involves ignoring evidence in one direction.

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- Intuitively, reject null of no effect when $|\overline{Y} \overline{X}|$ is large.



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 - Reject when $T_n \in C$.

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 - Medical diagnosis: false positive (type I) vs false negative (type II).

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- **Power function** of a test: probability of rejection as a function of *θ*:

 $\pi(\theta) = \mathbb{P}\left(\text{Reject } H_0 \mid \theta\right) = \mathbb{P}\left(T_n \in C \mid \theta\right)$

- **Hypotheticals!** if we knew θ , what is the probability of rejecting the null?
- The **power** of a test against an alternative $\theta_1 \in H_1$ is $\pi(\theta_1)$
- · We want to maximize power against alternative
- Size of a test is the probability of a Type I error:

$$\pi(\theta_0) = \mathbb{P}\left(\text{Reject } H_0 \mid \theta_0 \right)$$

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 - Still the dominant approach in the social sciences.



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 - Use the **quantile function**: $c = G_0^{-1}(1-\alpha)$
- If $G_0 \sim N(0, 1)$ and $\alpha = 0.05$, then $c = \Phi^{-1}(0.95) = 1.645$
 - Reject null if $T_n > 1.645$, fail to reject if $T_n \le 1.645$

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 - + For $\mathit{G}_{0}\sim\mathcal{N}(0,1)$ and $\alpha=$ 0.05, then c= 1.96

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• $|T_n| = 18.343 > 1.96 \rightsquigarrow \text{REJECT!}$

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- Size of the test converges to the **nominal size** as $n \to \infty$:

$$\mathbb{P}(|T_n| > z_{\alpha/2} \mid \theta_0) \to \alpha$$

3/ p-values

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$$p = \begin{cases} 1 - G_0(T_n) & \text{if one-sided} \\ 2(1 - G_0(|T_n|)) & \text{if two-sided} \end{cases}$$

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- p-hacking controversy: not about p-values per se, but about significance cutoffs

4/ Power Analyses

TABLE 2.	Effects of Four Mail Treatments on Voter Turnout in the August 2006 Primary
Election	• •

	Experimental Group				
	Control	Civic Duty	Hawthorne	Self	Neighbors
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$$T_n \stackrel{a}{\sim} \mathcal{N}\left(\frac{\theta_1}{\widehat{\mathsf{se}}[\hat{\theta}]}, 1\right) \qquad \rightsquigarrow \qquad \pi_n(\theta_1) = 1 - \Phi\left(c - \frac{\theta_1}{\widehat{\mathsf{se}}[\hat{\theta}]}\right)$$









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$$T_n = \frac{\overline{X}_n - \mu_0}{s_n / \sqrt{n}} \sim t_{n-1}$$

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- Null distribution is t so we use quantiles of t for critical values.
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 - · Asymptotically equivalent to using the normal, but more conservative

The shape of the t

